

Hillen's class, Oct. 15, 2024

### 3 Parameter Estimation

AIC Akaike Information Criterion

[Chapter 7, de Vries et al. A Course in Math Bio, SIAM 2006.]

Data, Observations

model. ODE e.g. oncolytic virus model

parameters  $p = (A, B, C, D)$

$$p \in \mathbb{R}^N$$

We assume that the data are typical!

we need a quality measure

$\mathcal{L}(p | \text{data})$  likelihood

assume  $\mathcal{L}(p | \text{data}) = \Pr(\text{data} | p)$

Definition.  $p \in \mathbb{R}^n$   $\mathcal{L}: \mathbb{R}^n \rightarrow \mathbb{R}_+$

which maps the parameters  $\pi \in \mathbb{R}^n$  to the probability to find given data  $\mathcal{L}(\pi) = \mathbb{P}_r(\text{data}|\pi)$   
 $\mathcal{L}$  is called the likelihood.

The log likelihood  $\ln \mathcal{L}(\pi) = \mathcal{L}\mathcal{L}(\pi)$

A parameter  $\hat{\pi}$  that maximizes  $\mathcal{L}$  (or  $\mathcal{L}\mathcal{L}$ ) is a maximum likelihood estimator.

Example 1. (time to cell death)

$$\dot{p} = -\mu p$$

$p$  pop

$\mu$  parameter const.

$$p(t) = p(0) e^{-\mu t}$$

death probability  $d(t) = 1 - p(t)$

$$= 1 - p(0) e^{-\mu t}$$

probability density of death  $f(t) = \frac{d}{dt} d(t)$

$$d'(t) = p(0) \mu e^{-\mu t}$$

assume  $p(0) = 1$

$$f(t) = \mu e^{-\mu t}$$

←

i.e. death is exponentially distributed  
and  $\frac{1}{\mu}$  is the expected survival time.

Suppose  $n$  independent cells die at ages  
 $a_1, a_2, \dots, a_n$

$$\frac{1}{\mu} \sim a_i \quad \underline{\underline{\mu \sim \frac{1}{a_i}}}$$

How to estimate  $\mu$ ?

1) average  $\frac{1}{a_1}, \frac{1}{a_2}, \dots, \frac{1}{a_n}$ :  $\hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n \frac{1}{a_i}$

2) average  $a_1, a_2, \dots, a_n$ :  $\hat{\mu}_2 = \frac{1}{\bar{a}}$   $\bar{a} = \frac{1}{n} \sum_{i=1}^n a_i$   
 $\hat{\mu}_1 \neq \hat{\mu}_2$  in general.  $\hat{\mu}_2$   $\bar{a}$

Likelihood  $\underline{\underline{L(\mu) = f(a_1) f(a_2) f(a_3) \dots f(a_n)}}$

$$= \mu e^{-\mu a_1} \mu e^{-\mu a_2} \dots \mu e^{-\mu a_n}$$

$$= \mu^n \exp\left(-\mu \sum_{i=1}^n a_i\right)$$

$L(\mu) = n \ln \mu - \mu \sum_{i=1}^n a_i$

 ←

$$\frac{d}{d\mu} \mathcal{L}(\mu) = \frac{n}{\mu} - \sum_{i=1}^n a_i = 0$$

$$\Rightarrow \frac{1}{\hat{\mu}} = \frac{1}{n} \sum a_i$$

$$\hat{\mu} = \left( \frac{1}{n} \sum_{i=1}^n a_i \right)^{-1} = \hat{\mu}_2$$

$$\frac{d^2}{d\mu^2} \mathcal{L}(\mu) = -\frac{n}{\mu^2} < 0 \Rightarrow \hat{\mu}_2 \text{ maximum likelihood estimator!}$$

Example: death times 1, 15, 3, 16, 18, 11, 20, 14, 13 =  $a_i$

$$\bar{a} = \frac{1}{n} \sum a_i = 12.3$$

$$\hat{\mu}_2 = \frac{1}{\bar{a}} = 0.081$$

$$\max \mathcal{L}(\hat{\mu}) = -31.6$$

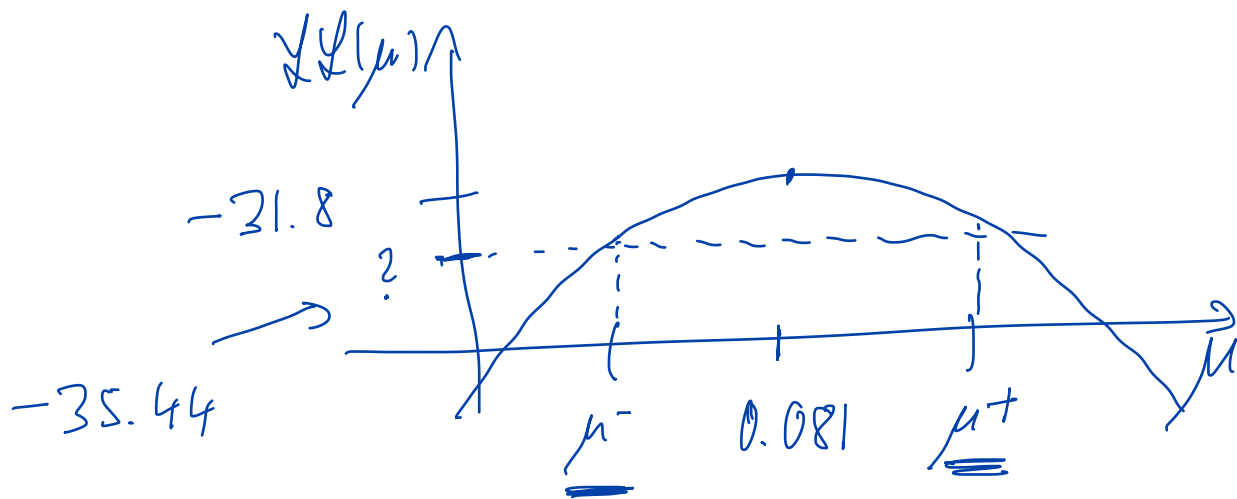
## Confidence intervals

Def: Given a confidence level  $\gamma$  (0.95 or 0.99) an interval  $[\rho^-, \rho^+]$  that covers the true parameter  $\rho$  with probability  $\gamma$  is called

a confidence interval.

Example: Cell death

$$\max \mathcal{L}(\mu) = -31.6$$



In most cases we assume  $\mathcal{L}(\mu)$  is  $\chi^2_1$  distributed. (one degree of freedom).

For example  $\gamma = 0.95$  tolerance  $\alpha = 1 - \gamma = 0.05$

$$\chi^2 \text{ table: } \chi^2_{1, \gamma}(0.05) = \underline{\underline{3.84}}$$

Now subtract 3.84 from  $\max \mathcal{L}(\mu)$

$$-31.6 - 3.84 = -35.44$$

Find  $\mu^-$ ,  $\mu^+$  from solving

$$\begin{aligned} \mathcal{L}(\mu^-) = \mathcal{L}(\mu^+) &= -35.44 \\ &= \max \mathcal{L}(\mu) - \chi^2_1(\alpha) \end{aligned}$$

in our example

$$n \ln \mu^+ - \mu^+ n \bar{a} = n \ln \mu^- - \mu^- n \bar{a} \\ = -35.44$$

$$\Rightarrow \mu^- = 0.027, \quad \mu^+ = 0.18$$

The 95% confidence interval is

$$[0.027, 0.18]$$

$$\text{and } \hat{\mu}_2^A = 0.081$$

## Least squares

at time points  $t_1, t_2, \dots, t_n$

measurements  $x_1, x_2, \dots, x_n$

which are fitted by a model  $f(t, p)$   
with parameters  $p$ .

$$x_i \approx \underline{f(t_i, p)} \quad i = 1, \dots, n$$

We assume

- $x_i$  independent
- normally distributed ←
- with expectation  $\mu = f(t_i, p)$
- and variance  $\sigma$  is unknown but const.

$$x_i \sim N(f(t_i, p), \sigma^2)$$

Statistical model ←  
ODE model

$$\begin{aligned} \mathcal{L}(\mu, \sigma^2) &= P(X_i = x_i, i=1, \dots, n \mid (\mu, \sigma)) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(x_i - f(t_i, p))^2}{2\sigma^2}} \end{aligned}$$

$$\mathcal{L}\mathcal{L}(\mu, \sigma^2) = - \sum_{i=1}^n \ln(\sqrt{2\pi} \sigma) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - f(t_i, p))^2$$

Annotations:  
-  $\ln(\sqrt{2\pi} \sigma)$  is crossed out with a red X.  
- Red arrows point from  $f(t_i, p)$  to "ODE" and "stat."  
- Red arrows point from the sum term to "minus".

To maximize  $\mathcal{L}\mathcal{L}(\mu, \sigma^2)$  we minimize  
 $\sum_{i=1}^n (x_i - f(t_i, p))^2$  : least square error, ☺

If variance varies

$$X_i \sim \mathcal{N}(f(t_i; \rho); \sigma_i^2)$$

Weighted least squares error

$$\min \left( \sum_{i=1}^n \frac{(X_i - f(t_i; \rho))^2}{\sigma_i^2} \right)$$

Growth  
curves

